

Hannes Strass

Faculty of Computer Science, Institute of Artificial Intelligence, Computational Logic Group

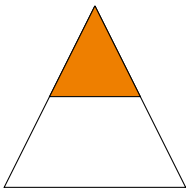
Games with Missing Information: Modelling

Lecture 7, 1st June 2026 // Algorithmic Game Theory, SS 2026

Previously ...

- **Monte Carlo Tree Search** uses random playouts to evaluate moves and keeps statistics on which moves led to which payoffs how many times.
- A **selection policy** balances **exploitation** and **exploration**.
- **UCT** is an effective selection policy that applies UCB1 to trees.
- A **playout policy** steers playout simulations towards realistic play.
- MCTS and deep reinforcement learning led to expert-level Go programs.

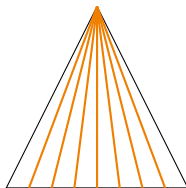
wide, but shallow



Type A

Alpha-Beta Tree Search

narrow, but deep



Type B

Monte Carlo Tree Search

Overview

Example: The Monty Hall Problem

Extensive-Form Games

Behaviour Strategies

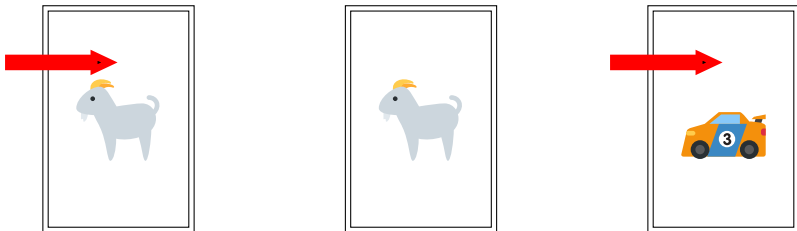
Kuhn's Theorem

Motivation: Missing Information

- So far, we have considered games with **perfect** information:
- In every state, all players know the full history of play so far, i.e. they know their (joint) position in the game tree.
- However, e.g. in card games, players typically do not know the cards of opponents.
- This form of incomplete knowledge can be formalised by sets of indistinguishable nodes in the game tree, typically called **information sets**.
- In this context, we also add another element to games: **chance**.
- This is modelled via **moves by nature** and can be used to formalise dealing cards or throwing dice.
- We will see that this also allows us to model games with **incomplete** information, where e.g. some of the payoffs may be uncertain.
- In principle, however, **chance** and **imperfect information** are unrelated and we could model either without the other.

Example: The Monty Hall Problem

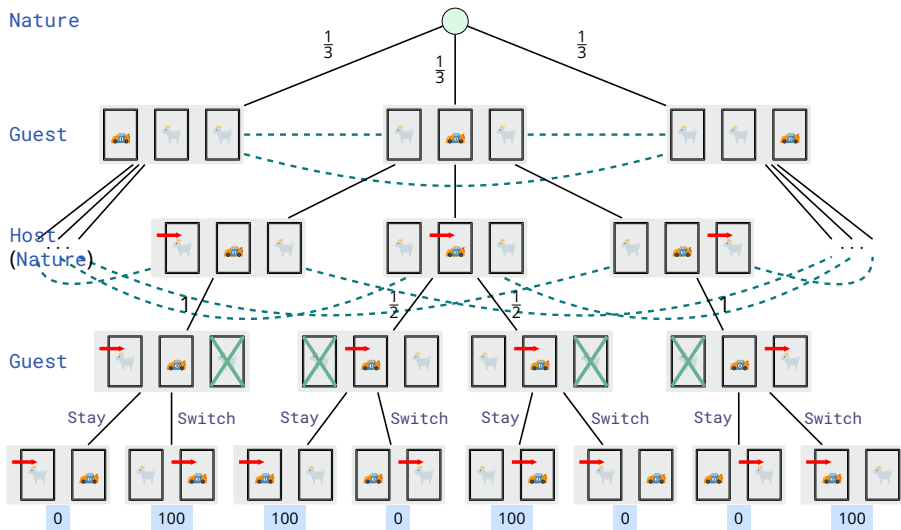
Example: The Monty Hall Problem



The Monty Hall Problem

A game show participant (**Guest**) is shown three doors behind which there are prizes. Behind one door, there is an expensive car, behind each of the other doors there is a goat. (The participant prefers the car over a goat.) The **Guest** is asked to **Choose** one of the doors. The game show **Host** (the other player) now opens one of the remaining doors that has a goat behind it. The **Guest** then gets their final move: **Stay** with the door they initially picked, or **Switch** to the other door. What should the participant do?

The Monty Hall Problem: Game Tree Sketch



The Monty Hall Problem: Analysis

- Each of the possible states s_1, s_2, s_3 after **Nature's** move has probability $\frac{1}{3}$.
- For each of these states, the ensuing game is symmetric.
- If **Guest** chooses their door uniformly at random, then:
 - With probability $\frac{1}{3}$, their initial guess is correct; thus **Switch** has a payoff of 0.
 - With probability $\frac{2}{3}$, their initial guess is wrong; thus **Switch** has a payoff of 100.
- Thus in each s_i , **Switch** has an expected payoff of $\frac{1}{3} \cdot 0 + \frac{2}{3} \cdot 100$.
- The overall payoff of **Switch** is thus

$$U_{\text{Guest}}(\text{Switch}) = 3 \cdot \frac{1}{3} \cdot \left(\frac{1}{3} \cdot 0 + \frac{2}{3} \cdot 100 \right) = 66\frac{2}{3}$$

- Likewise, the overall payoff of **Stay** is obtained as

$$U_{\text{Guest}}(\text{Stay}) = \frac{1}{3} \cdot 100 = 33\frac{1}{3}$$

- Therefore, a rational player should always choose **Switch** over **Stay**.

Extensive-Form Games

Missing Information: Formalisation

Definition

An **extensive-form game** consists of the following:

1. A set $P = \{1, \dots, n\}$ of players, and possibly **Nature**.
2. An $n + 1$ -tuple $\mathbf{M} = (M_1, \dots, M_n, M_{\text{Nature}})$ of sets M_i of **moves** for all players.
3. A set H of **histories**, sequences of moves $m_j \in M_1 \cup \dots \cup M_n \cup M_{\text{Nature}}$.
4. A subset $Z \subseteq H$ of **terminal** histories.
5. A partition $\mathcal{J}_1 \dot{\cup} \dots \dot{\cup} \mathcal{J}_k = H \setminus Z$ of non-terminal histories into **information sets** such that for all $1 \leq j \leq k$, all $h_1, h_2 \in \mathcal{J}_j$ have the same legal moves.
6. A **player function** $p: \{1, \dots, k\} \rightarrow P \cup \{\text{Nature}\}$ (stating whose turn it is).
7. An n -tuple $\mathbf{u} = (u_1, \dots, u_n)$ of utility functions $u_i: Z \rightarrow \mathbb{R}$.

Starting with the **empty history** $[],$ in each history $h = [m_1, \dots, m_\ell] \in H \setminus Z,$ player $i = p(h)$ chooses a move $m \in M_i,$ leading to the history $[m_1, \dots, m_\ell, m].$ Each \mathcal{J}_j with $p(\mathcal{J}_j) = \text{Nature}$ has a probability distribution on possible moves.

Information Sets: Remarks

Intuition of information sets: The player (whose turn it is) does not have the information to distinguish between states in the set (but from other sets).

- $\mathcal{I} = \{\mathcal{I}_1, \dots, \mathcal{I}_k\}$ being a **partition** means that:
 - for all $1 \leq j \leq k$, we have $\mathcal{I}_j \neq \emptyset$,
 - $\mathcal{I}_1 \cup \dots \cup \mathcal{I}_k = H \setminus Z$, and
 - for all $1 \leq j, \ell \leq k$, we have $\mathcal{I}_j \cap \mathcal{I}_\ell = \emptyset$.
- Thus every $h \in H \setminus Z$ belongs to exactly one information set $\mathcal{I}_h \in \mathcal{I}$.
- For all $1 \leq j \leq k$ and $h \in \mathcal{I}_j$, we denote $p(h) := p(\mathcal{I}_j)$.
- \mathcal{I} can also be represented by an **equivalence relation** \sim^G , where for any $h_1, h_2 \in H$, we have $h_1 \sim^G h_2$ iff there is a $\mathcal{I}_j \in \mathcal{I}$ such that $h_1, h_2 \in \mathcal{I}_j$.
- We graphically represent \sim^G in game trees via dashed edges $-----$.

Battleship

The initial placement of ships is private to the players and can be modelled via information sets. Some information may later be disclosed through hits.

Information Sets: Example

The Monty Hall Problem

- The (true) initial state is represented by the information set $\mathcal{J}_0 = \{[]\}$.
- The (seemingly) initial state for **Guest** is given by the information set
$$\mathcal{J}_1 = \{[\text{Car1}], [\text{Car2}], [\text{Car3}]\}.$$
- For each possible (initial) choice of door for **Guest**, there is one set:

$$\mathcal{J}_{\text{Choose1}} = \{[\text{Car1}, \text{Choose1}], [\text{Car2}, \text{Choose1}], [\text{Car3}, \text{Choose1}]\}$$

$$\mathcal{J}_{\text{Choose2}} = \{[\text{Car1}, \text{Choose2}], [\text{Car2}, \text{Choose2}], [\text{Car3}, \text{Choose2}]\}$$

$$\mathcal{J}_{\text{Choose3}} = \{[\text{Car1}, \text{Choose3}], [\text{Car2}, \text{Choose3}], [\text{Car3}, \text{Choose3}]\}$$

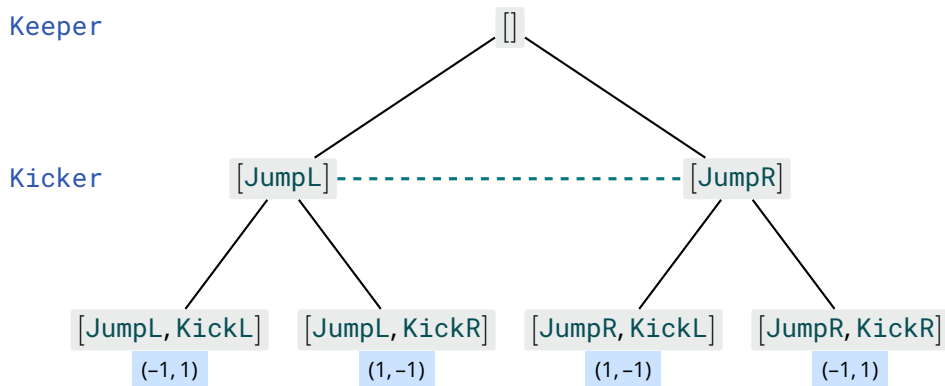
- Some information is disclosed by the host opening a door:

$$\mathcal{J}_{[\text{Choose1}, \text{open2}]} = \{[\text{Car1}, \text{Choose1}], [\text{Car3}, \text{Choose1}]\}$$

Strategic Games and Imperfect Information

- Uncertainty induced by simultaneous moves can be modelled in extensive-form games (that seem to be sequential by definition).
- **Main idea:** Sequentialise moves, model uncertainty in information sets.

Example: Recall the game penalties. One extensive-form variant is:



Chance Nodes (Moves by Nature)

Intuition of chance nodes: Something happens that is controlled by an entity with no strategic interest in the game's outcome.

Examples

- In card games, **Nature** controls the dealer's shuffling the cards.
- In games involving dice, **Nature** controls the dice throws.
- Probability distributions model uncertainty about effects of such actions.
- We typically use uniform distributions over possible atomic results.
- In the remainder of the course, we use **conditional probabilities** ...
- ...and how they relate to marginal probabilities via **Bayes' Theorem**.
- The relevant definitions and results are in the appendix.

Behaviour Strategies

From Extensive Form to Normal Form

Definition

Let $G = (P, \mathbf{M}, H, Z, \{\mathcal{I}_1, \dots, \mathcal{I}_k\}, p, \mathbf{u})$ be a finite extensive-form game.

1. For $i \in P$, we denote $p^{-1}(i) = \{\mathcal{I}_j \in \mathcal{I} \mid p(\mathcal{I}_j) = i\}$.
2. The **possible moves** of player $i \in P$ at information set $\mathcal{I}_j \in p^{-1}(i)$ are:

$$M_i(\mathcal{I}_j) = \{m \in M_i \mid [h; m] \in H \text{ for some } h \in \mathcal{I}_j\}$$

3. A **pure strategy** for player $i \in P$ is a function

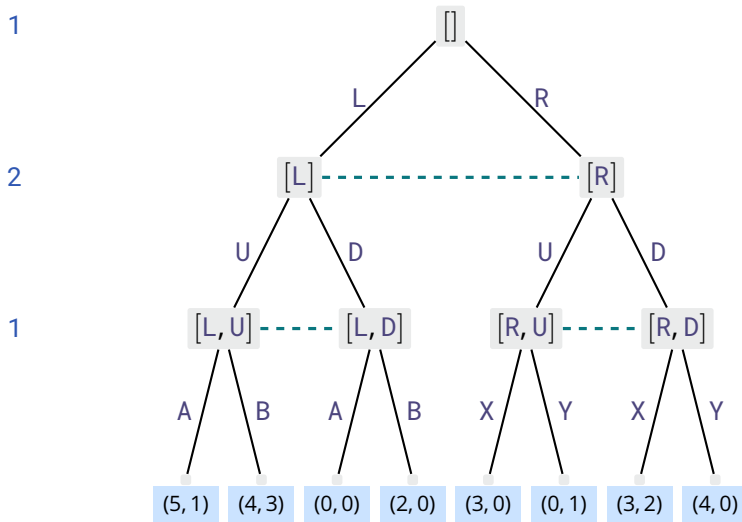
$$s_i: p^{-1}(i) \rightarrow M_i \quad \text{where} \quad s_i(\mathcal{I}_j) \in M_i(\mathcal{I}_j) \text{ for all } \mathcal{I}_j \in p^{-1}(i)$$

The **associated normal-form game** $G' = (P, \mathbf{S}, \mathbf{u}')$ is as follows:

- $\mathbf{S} = (S_1, \dots, S_n)$ with each S_i the set of all pure strategies for player i .
- $\mathbf{u}' = (u'_1, \dots, u'_n)$ with each u'_i mapping profile $\mathbf{s} = (s_1, \dots, s_n)$ to $u_i(\langle\langle \mathbf{s}, [] \rangle\rangle)$, where $\langle\langle \mathbf{s}, [] \rangle\rangle$ is the terminal history obtained from playing as specified by \mathbf{s} .

As usual, strategies can be reduced; pure strategies can be mixed.

Extensive to Normal Form: Example



(1, 2)	U	D
LA*	(5, 1)	(0, 0)
LB*	(4, 3)	(2, 0)
R*X	(3, 0)	(3, 2)
R*Y	(0, 1)	(4, 0)

Strategies and the Normal Form: Remarks

- Playing a mixed strategy $\pi: S_i \rightarrow [0, 1]$ in an extensive-form game can be understood as randomly choosing a **pure** strategy $s_i \sim \pi$ and then strictly following this pure strategy s_i during gameplay.
- Decisions in different parts of the game tree are determined by the chosen s_i and thus **not independent**.
- For an extensive-form game $G = (P, \mathbf{M}, H, Z, \mathcal{I}, p, \mathbf{u})$, player $i \in P$ may have

$$|M_i|^{|\rho^{-1}(i)|}$$

different pure strategies (in the normal-form game) in the worst case.

- Practically, we are however mainly interested in **local** decision making.
- In a **behaviour strategy** σ , a player has a probability distribution $\sigma(\mathcal{I}_j)$ in each information set \mathcal{I}_j and randomly chooses a **move** $m \sim \sigma(\mathcal{I}_j)$ when in \mathcal{I}_j .
- Probability distributions at different information sets are **independent**.

Behaviour Strategies

Definition

Let G be an extensive-form game with players P and information sets \mathcal{I} . A **behaviour strategy** for player $i \in P$ is a function σ_i that assigns a probability distribution over possible moves to each of player i 's information sets.

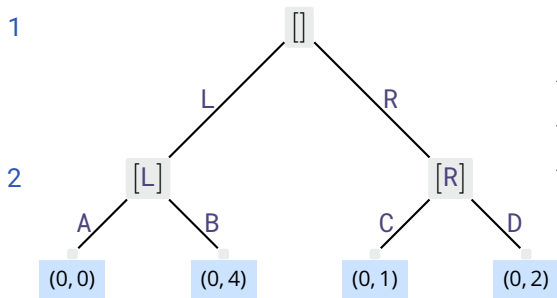
- $\sigma_i(\mathcal{I}_j)(m_k)$ denotes the probability that player i will make move m_k at information set \mathcal{I}_j . For readability, we will write this as $\sigma_i(m_k | \mathcal{I}_j)$.
- Every behaviour strategy σ of player $i \in P$ induces a mixed strategy π_σ :

$$\pi_\sigma := \left\{ s_i \mapsto \prod_{\mathcal{I}_j \in p^{-1}(i)} \sigma(s_i(\mathcal{I}_j) | \mathcal{I}_j) \mid s_i \in S_i \right\}$$

- In particular, a pure strategy s_i with $s_i(\mathcal{I}_j) = m_k$ can be seen as a behaviour strategy σ_i with $\sigma_i(m_k | \mathcal{I}_j) = 1$ and $\sigma_i(m_\ell | \mathcal{I}_j) = 0$ for $m_\ell \in M_i(\mathcal{I}_j)$, $\ell \neq k$.

Behaviour vs. Mixed Strategies: Examples

Consider the following extensive-form game G_5 and its normal form:



(1, 2)	AC	AD	BC	BD
L	(0, 0)	(0, 0)	(0, 4)	(0, 4)
R	(0, 1)	(0, 2)	(0, 1)	(0, 2)

- A mixed strategy for player 2 is $\pi = \left(\frac{1}{2}, 0, 0, \frac{1}{2}\right) \triangleq \left\{AC \mapsto \frac{1}{2}, BD \mapsto \frac{1}{2}, \dots\right\}$.
- The “closest” behaviour strategy σ assigns $\sigma([L]) = \left\{A \mapsto \frac{1}{2}, B \mapsto \frac{1}{2}\right\}$ and $\sigma([R]) = \left\{C \mapsto \frac{1}{2}, D \mapsto \frac{1}{2}\right\}$, inducing mixed strategy $\pi_\sigma = \left(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4}\right) \neq \pi$.

Kuhn's Theorem

Idea and Intuition

Do we always need the “full power” of mixed strategies to analyse extensive-form games?

- Representing a mixed strategy requires up to $|M_i|^{|\rho^{-1}(i)|}$ numbers.
- Representing a behaviour strategy requires up to $|M_i| \cdot |\rho^{-1}(i)|$ numbers.

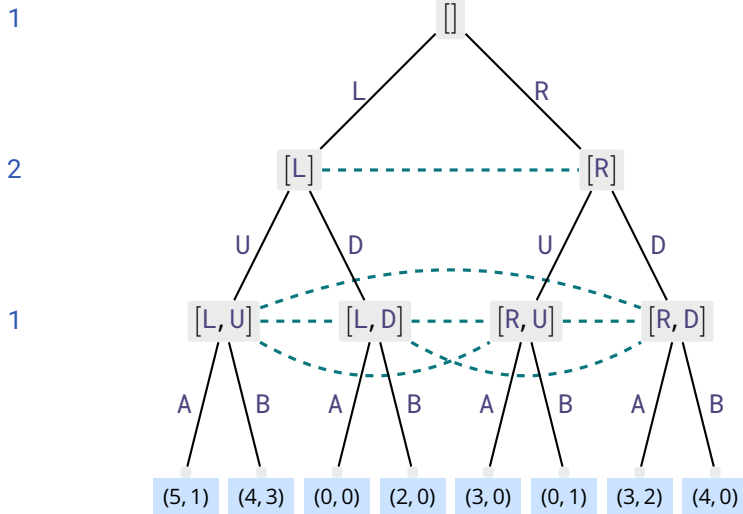
Question

Can mixed strategies always be equivalently written as behaviour strategies?

What does “equivalent” even mean in this context?

↔ Intuitively: When all histories are reached with the same probability.
(Formal definition later.)

Equivalent Behaviour Strategies: Example



(1, 2)	U	D
LA	(5, 1)	(0, 0)
LB	(4, 3)	(2, 0)
RA	(3, 0)	(3, 2)
RB	(0, 1)	(4, 0)

Player 1's mixed strategy $(0, \frac{2}{5}, \frac{3}{5}, 0)$ has no equivalent behaviour strategy.

Perfect Recall

Definition

Let $G = (P, \mathbf{M}, H, \mathcal{J}, p, \mathbf{u})$ be an extensive-form game.

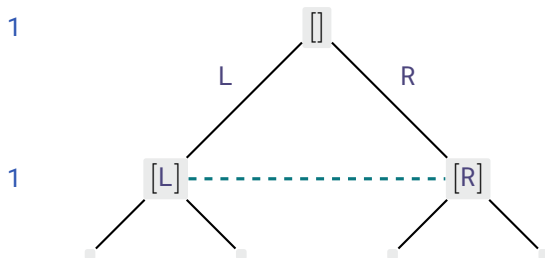
- For every player $i \in P$ and history $h \in H$, define the sequence h_i of pairs (\mathcal{J}_j, m) for $\mathcal{J}_j \in \mathcal{J}$ and $m \in M_j$ by induction:

$$[]_i := [] \quad \text{and} \quad [h; m]_i := \begin{cases} [h_i; (\mathcal{J}_h, m)] & \text{if } p(h) = i, \\ h_i & \text{otherwise.} \end{cases}$$

where for $h = [m_1, \dots, m_\ell]$ we denote $[h; m] = [m_1, \dots, m_\ell, m]$ as usual.

- Player $i \in P$ has **perfect recall** in G iff for all $\mathcal{J}_j \in \mathcal{J}$, for every $h, h' \in \mathcal{J}_j$, it holds that $h_i = h'_i$.
 - G has **perfect recall** iff every player $i \in P$ has perfect recall in G .
- h_i extracts all decision points and decisions (of player i) from history h .
 - $h_i = h'_i$ means that i **made the same moves in the same information sets**.
 - With perfect recall, players remember their trajectory through the game.

Perfect Recall: Examples (1)

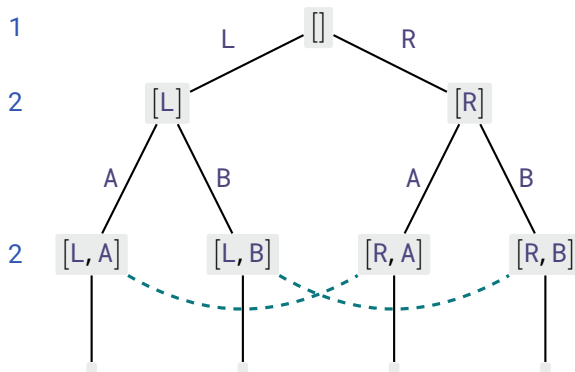


This game does not have perfect recall:

- Denote $\mathcal{J}_0 = \{[]\}$ and $\mathcal{J}_1 = \{[L], [R]\}$.
- We have $[L] \in \mathcal{J}_1$ and $[R] \in \mathcal{J}_1$, but:

$$[L]_1 = [(\mathcal{J}_0, L)] \neq [(\mathcal{J}_0, R)] = [R]_1$$

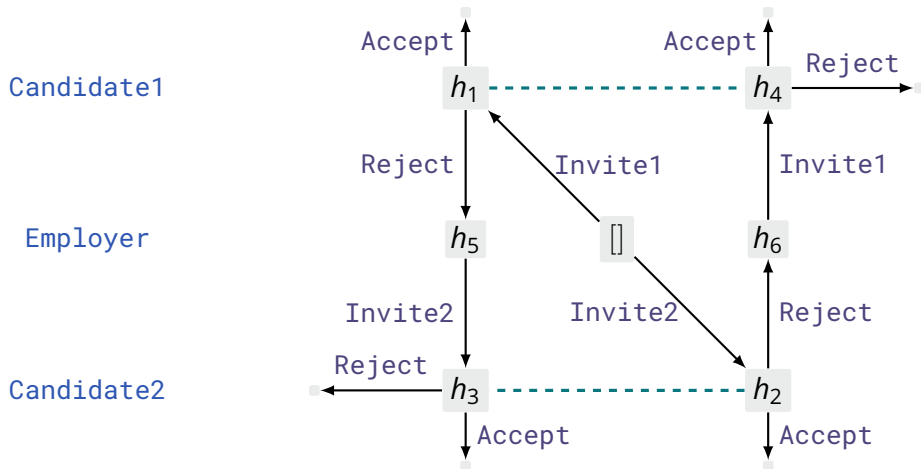
Perfect Recall: Examples (2)



This game does not have perfect recall:

- Denote $\mathcal{J}_0 = \{[\]\}$, $\mathcal{J}_1 = \{[L]\}$, $\mathcal{J}_2 = \{[R]\}$, $\mathcal{J}_3 = \{[L, A], [R, A]\}$, $\mathcal{J}_4 = \{[L, B], [R, B]\}$.
- Then $[L, A]_2 = [(\mathcal{J}_1, A)] \neq [(\mathcal{J}_2, A)] = [R, A]_2$.

Perfect Recall: Examples (3)



This game has perfect recall: e.g. $h \in \{h_2, h_3\}$ implies $h_{\text{Candidate2}} = \{\}$.

Realisation Equivalence

Definition

Let $G = (P, \mathbf{M}, H, Z, \mathcal{J}, p, \mathbf{u})$ be an extensive-form game and $i \in P$.

1. For every history $h \in H$ and mixed strategy $\pi_i: S_i \rightarrow [0, 1]$, the probability of i moving as in h_i when playing according to π is $P(h_i | \pi)$ with

$$P([(J_1, m_1), \dots, (J_\ell, m_\ell)] | \pi) := \sum_{\substack{s_i \in S_i, \\ s_i(J_1) = m_1, \dots, s_i(J_\ell) = m_\ell}} \pi(s_i)$$

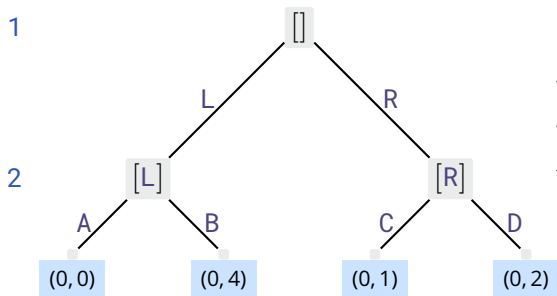
2. For behaviour strategy σ_i , $P([(J_1, m_1), \dots, (J_\ell, m_\ell)] | \sigma_i) := \prod_{1 \leq j \leq \ell} \sigma_i(m_j | J_j)$.

3. Mixed strategy π and behaviour strategy σ for player $i \in P$ are **realisation-equivalent** iff for all $h \in H$: $P(h_i | \pi) = P(h_i | \sigma)$.

In particular $P([\] | \pi) = \sum_{s_i \in S_i} \pi(s_i) = 1 = P([\] | \sigma)$.

Realisation Equivalence: Example & Remark

Consider again extensive-form game G_5 :



(1, 2)	AC	AD	BC	BD
L	(0, 0)	(0, 0)	(0, 4)	(0, 4)
R	(0, 1)	(0, 2)	(0, 1)	(0, 2)

- Player 2's mixed strategy $\pi = \left(\frac{1}{2}, 0, 0, \frac{1}{2}\right)$ and behaviour strategy σ with $\sigma(A | [L]) = \sigma(B | [L]) = \frac{1}{2}$ and $\sigma(C | [R]) = \sigma(D | [R]) = \frac{1}{2}$ are realisation-equivalent:

$$P(\{[L]\}, A | \pi) = \frac{1}{2} = P(\{[L]\}, A | \sigma) \text{ and similarly for } \{[R]\}.$$

- Generally, since in particular terminal histories are reached with the same probabilities, realisation-equivalent strategies obtain equal game values.

Kuhn's Theorem

Theorem [Kuhn, 1953]

Let G be an extensive-form game with perfect recall. For every mixed strategy π for i , there exists a realisation-equivalent behaviour strategy σ .

Proof.

- Since G has perfect recall, for any $\mathcal{J}_j \in p^{-1}(i)$ we have that $h', h'' \in \mathcal{J}_j$ satisfy $h'_i = h''_i$ and thus any $h \in \mathcal{J}_j$ uniquely defines $P(\mathcal{J}_j | \pi) := P(h_i | \pi)$.
- For $\mathcal{J}_h \in p^{-1}(i)$ with $P(\mathcal{J}_h | \pi) > 0$ and $m \in M_i(\mathcal{J}_h)$, set $\sigma(m | \mathcal{J}_h) := \frac{P([h_i; (\mathcal{J}_h, m)] | \pi)}{P(h_i | \pi)}$.
- To show realisation equivalence, consider any $h_i = [(\mathcal{J}_1, m_1), \dots, (\mathcal{J}_\ell, m_\ell)]$:

$$\begin{aligned} P(h_i | \sigma) &= \sigma(m_1 | \mathcal{J}_1) \cdot \sigma(m_2 | \mathcal{J}_2) \cdot \dots \cdot \sigma(m_\ell | \mathcal{J}_\ell) \\ &= \frac{P([(\mathcal{J}_1, m_1)] | \pi)}{P(\square | \pi)} \cdot \frac{P([(\mathcal{J}_1, m_1), (\mathcal{J}_2, m_2)] | \pi)}{P([(\mathcal{J}_1, m_1)] | \pi)} \cdot \dots \cdot \frac{P([(\mathcal{J}_1, m_1), \dots, (\mathcal{J}_\ell, m_\ell)] | \pi)}{P([(\mathcal{J}_1, m_1), \dots, (\mathcal{J}_{\ell-1}, m_{\ell-1})] | \pi)} \\ &= \frac{P([(\mathcal{J}_1, m_1), \dots, (\mathcal{J}_\ell, m_\ell)] | \pi)}{P(\square | \pi)} = P(h_i | \pi) \end{aligned}$$

□

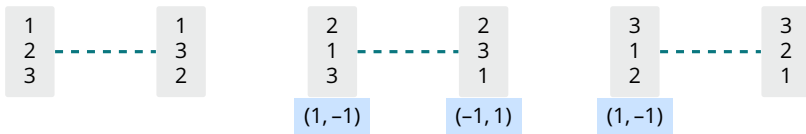
Example: Simplified Poker

Binmore's Simplified Poker

- Two players, **Ann** and **Bob**, each put \$1 into a jackpot.
- They then draw one card from a deck of three cards: {1, 2, 3}.
- **Ann** can either **check** (pass on), or **raise** (put another \$1 into the jackpot).
- Next, **Bob** responds:
 - If **Ann** has **checked**, then **Bob** must **call**, that is, a **showdown** happens: Both players show their cards and the player with the higher (number) card receives the jackpot.
 - If **Ann** has **raised**, then **Bob** can decide between **fold** (withdraw from the game and let **Ann** get the jackpot) or **call** (put another \$1 into the jackpot and then have a showdown).

Simplified Poker: Preliminary Analysis

Nature shuffles and deals the cards. There are six possible outcomes:



- If Ann draws a 3, she will **raise**; if Bob draws a 1, he will **fold**.
- If Bob draws a 3, he will **call**; if Ann draws a 2, she will **check**:
Were she to **raise**, she would lose 2 if Bob has a 3 (as he would **call**), but still only win 1 if Bob has a 1 (as he would **fold** then).

What happens in the two remaining cases?

1. Should Ann **raise** (i.e. bluff) if she has a 1?
2. Should Bob **call** (the bluff) if he has a 2?

Conclusion

Summary

- In **complete information** games, players know the rules, possible outcomes and each other's preferences over outcomes.
- In **perfect information** games, all players know all previous moves.
- An **extensive-form** game may have incomplete or imperfect information.
- Uncertainty of players (due to missing information) can be modelled by **information sets** and **chance nodes** (moves by *Nature*).
- A **behaviour strategy** assigns move probabilities to information sets.
- In a game with **perfect recall**, players remember their own previous moves and the information sets in which they played them.
- **Kuhn's Theorem**: Considering only behaviour strategies "is enough" for playing games with perfect recall.

Goat and Car graphics: Twemoji, Copyright 2020 Twitter, Inc and other contributors (CC-BY 4.0)

Appendix: Probabilities

Recall

- A **probability space** is a finite set $\mathcal{E} = \{e_1, \dots, e_k\}$ of **atomic events**.
- A **probability distribution** is a mapping $P: \mathcal{E} \rightarrow [0, 1]$, where atomic event e_i occurs with probability $P(e_i)$ and we have $\sum_{i=1}^k P(e_i) = 1$.
- An **event** $E \subseteq \mathcal{E}$ has (total) probability $P(E) = \sum_{e \in E} P(e)$.
- For all events $A, B \subseteq \mathcal{E}$ we have the following:
 1. $0 \leq P(A) \leq 1$ with $P(\emptyset) = 0$ and $P(\mathcal{E}) = 1$.
 2. $P(\bar{A}) = 1 - P(A)$ where $\bar{A} := \mathcal{E} \setminus A$ is the event **complementary** to A .
 3. $P(A \cup B) = P(A) + P(B) - P(A \cap B)$.

Example

If all events $e_i \in \mathcal{E}$ have the same probability $\frac{1}{|\mathcal{E}|}$, we have a **uniform distribution**.

Conditional Probabilities

Definition

Let A and B be events with $P(B) > 0$.

1. The **conditional probability** for A to occur under the condition of B occurring is

$$P(A|B) := \frac{P(A \cap B)}{P(B)}$$

2. Events A and B are **independent** iff

$$P(A \cap B) = P(A) \cdot P(B)$$

That events A and B are independent is equivalently characterised by each of:

- $P(A|B) = P(A)$
- $P(A|\bar{B}) = P(A)$
- $P(B|A) = P(B)$
- $P(B|\bar{A}) = P(B)$

Bayes' Theorem

Theorem (Bayes)

1. If A and B are two events with $P(A) > 0$ and $P(B) > 0$, then

$$P(A) \cdot P(B|A) = P(B) \cdot P(A|B)$$

2. If A and B_1, B_2, \dots, B_ℓ are events with $P(A) > 0$ and $P(B_i) > 0$ for all $1 \leq i \leq \ell$, where $\bigcup_{i=1}^{\ell} B_i = \mathcal{E}$ is a partition of \mathcal{E} , then for every $1 \leq i \leq \ell$:

$$P(B_i|A) = \frac{P(A|B_i) \cdot P(B_i)}{\sum_{j=1}^{\ell} (P(A|B_j) \cdot P(B_j))} = \frac{P(A|B_i) \cdot P(B_i)}{P(A)}$$

In the second item of the theorem, the **law of total probability** is used:

$$P(A) = \sum_{j=1}^{\ell} P(B_j \cap A) = \sum_{j=1}^{\ell} (P(A|B_j) \cdot P(B_j))$$

Note that $P(A) = P(\mathcal{E} \cap A) = P((\bigcup_{j=1}^{\ell} B_j) \cap A) = P(\bigcup_{j=1}^{\ell} (B_j \cap A)) = \sum_{j=1}^{\ell} P(B_j \cap A)$.